



7. Loss systems

lect07.ppt

S-38.145 - Introduction to Teletraffic Theory - Fall 2000

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7. Loss systems

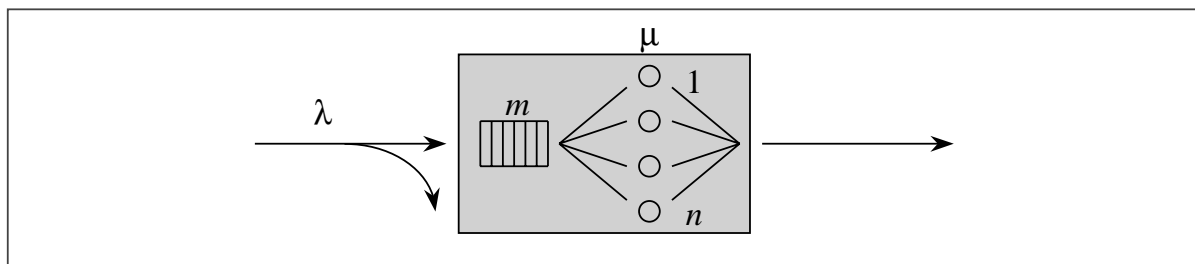
Contents

- Refresher: Simple teletraffic model
- Poisson model (∞ customers, ∞ servers)
- Erlang model (∞ customers, $n < \infty$ servers)
- Binomial model ($k < \infty$ customers, $n = k$ servers)
- Engset model ($k < \infty$ customers, $n < k$ servers)

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Simple teletraffic model

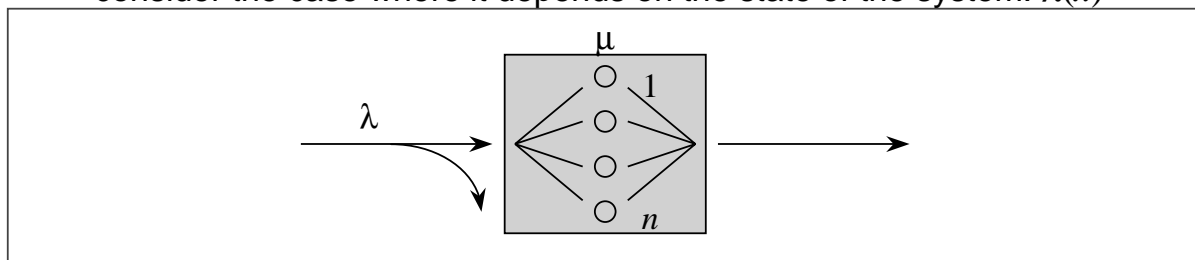
- **Customers arrive** at rate λ (customers per time unit)
 - $1/\lambda$ = average inter-arrival time
- Customers are **served** by n parallel **servers**
- When busy, a server serves at rate μ (customers per time unit)
 - $1/\mu$ = average service time of a customer
- There are m **waiting** places
- It is assumed that blocked customers (arriving in a full system) are lost



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Pure loss system

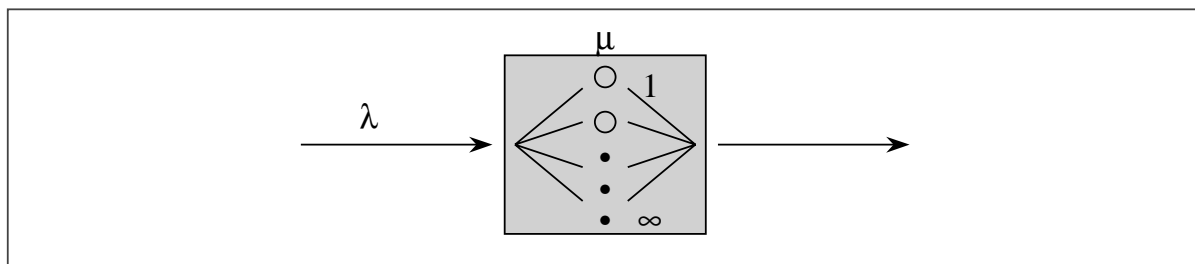
- No waiting places ($m = 0$)
 - If the system is full (with all n servers occupied) when a customer arrives, she is not served at all but lost
 - Some customers are lost
- From the customer's point of view,
 - it is interesting to know e.g. the blocking probability
- Note: In addition to the case where the arrival rate λ is constant, we will consider the case where it depends on the state of the system: $\lambda(x)$



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Infinite system

- Infinite number of servers ($n = \infty$)
 - No customers are lost or even have to wait before getting served
- Note: Also here, in addition to the case where the arrival rate λ is constant, we will consider the case where it depends on the state of the system: $\lambda(x)$



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Blocking

- In a loss system some calls are lost
 - a call is lost if all n channels are occupied when the call arrives
 - the term **blocking** refers to this event
- There are (at least) two different types of blocking quantities:
 - **Call blocking** B_c = probability that an arriving call finds all n channels occupied = the fraction of calls that are lost
 - **Time blocking** B_t = probability that all n channels are occupied at an arbitrary time = the fraction of time that all n channels are occupied
- The two blocking quantities are not necessarily equal
 - If calls arrive according to a Poisson process, then $B_c = B_t$
- Call blocking is a better measure for the quality of service experienced by the subscribers but, typically, time blocking is easier to calculate

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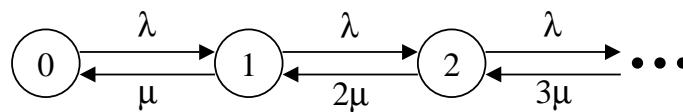
Poisson model (M/M/ ∞)

- **Definition: Poisson model** is the following simple teletraffic model:
 - Infinite number of independent customers ($k = \infty$)
 - Interarrival times are IID and exponentially distributed with mean $1/\lambda$
 - so, customers arrive according to a Poisson process with intensity λ
 - Infinite number of servers ($n = \infty$)
 - Service times are IID and exponentially distributed with mean $1/\mu$
 - No waiting places ($m = 0$)
- Poisson model:
 - Using Kendall's notation, this is an M/M/ ∞ queue
 - Infinite system, and, thus, **lossless**
- Notation:
 - $a = \lambda/\mu =$ traffic intensity

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State transition diagram

- Let $X(t)$ denote the number of customers in the system at time t
 - Assume that $X(t) = i$ at some time t , and consider what happens during a short time interval $(t, t+h]$:
 - with prob. $\lambda h + o(h)$,
a new customer arrives (state transition $i \rightarrow i+1$)
 - if $i > 0$, then, with prob. $i\mu h + o(h)$,
a customer leaves the system (state transition $i \rightarrow i-1$)
- Process $X(t)$ is clearly a Markov process with state transition diagram



- Note that process $X(t)$ is an irreducible birth-death process with an infinite state space $S = \{0, 1, 2, \dots\}$

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Equilibrium distribution (1)

- Local balance equations (LBE):

$$\pi_i \lambda = \pi_{i+1} (i+1) \mu \quad (\text{LBE})$$

$$\Rightarrow \pi_{i+1} = \frac{\lambda}{(i+1)\mu} \pi_i = \frac{a}{i+1} \pi_i$$

$$\Rightarrow \pi_i = \frac{a^i}{i!} \pi_0, \quad i = 0, 1, 2, \dots$$

- Normalizing condition (N):

$$\sum_{i=0}^{\infty} \pi_i = \pi_0 \sum_{i=0}^{\infty} \frac{a^i}{i!} = 1 \quad (\text{N})$$

$$\Rightarrow \pi_0 = \left(\sum_{i=0}^{\infty} \frac{a^i}{i!} \right)^{-1} = (e^a)^{-1} = e^{-a}$$

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Equilibrium distribution (2)

- Thus, the equilibrium distribution is a **Poisson distribution**:

$$X \sim \text{Poisson}(a)$$

$$P\{X = i\} = \pi_i = \frac{a^i}{i!} e^{-a}, \quad i = 0, 1, 2, \dots$$

$$E[X] = a, \quad D^2[X] = a$$

- Remark (insensitivity):
 - The result is **insensitive** to the service time distribution, that is: it is valid for **any** service time distribution with mean $1/\mu$
 - So, instead of the $M/M/\infty$ model, we can consider, as well, the more general $M/G/\infty$ model

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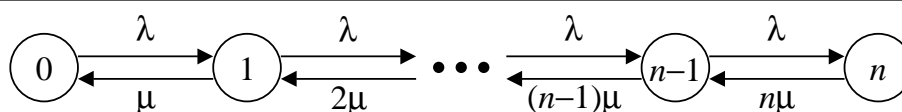
Erlang model (M/M/n/n)

- **Definition: Erlang model** is the following simple teletraffic model:
 - Infinite number of independent customers ($k = \infty$)
 - Interarrival times are IID and exponentially distributed with mean $1/\lambda$
 - so, customers arrive according to a Poisson process with intensity λ
 - Finite number of servers ($n < \infty$)
 - Service times are IID and exponentially distributed with mean $1/\mu$
 - No waiting places ($m = 0$)
- Erlang model:
 - Using Kendall's notation, this is an M/M/n/n queue
 - Pure loss system, and, thus, **lossy**
- Notation:
 - $a = \lambda/\mu =$ traffic intensity

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State transition diagram

- Let $X(t)$ denote the number of customers in the system at time t
 - Assume that $X(t) = i$ at some time t , and consider what happens during a short time interval $(t, t+h]$:
 - with prob. $\lambda h + o(h)$,
a new customer arrives (state transition $i \rightarrow i+1$)
 - with prob. $i\mu h + o(h)$,
a customer leaves the system (state transition $i \rightarrow i-1$)
- Process $X(t)$ is clearly a Markov process with state transition diagram



- Note that process $X(t)$ is an irreducible birth-death process with a finite state space $S = \{0, 1, 2, \dots, n\}$

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Equilibrium distribution (1)

- Local balance equations (LBE):

$$\pi_i \lambda = \pi_{i+1} (i+1) \mu \quad (\text{LBE})$$

$$\Rightarrow \pi_{i+1} = \frac{\lambda}{(i+1)\mu} \pi_i = \frac{a}{i+1} \pi_i$$

$$\Rightarrow \pi_i = \frac{a^i}{i!} \pi_0, \quad i = 0, 1, \dots, n$$

- Normalizing condition (N):

$$\sum_{i=0}^n \pi_i = \pi_0 \sum_{i=0}^n \frac{a^i}{i!} = 1 \quad (\text{N})$$

$$\Rightarrow \pi_0 = \left(\sum_{i=0}^n \frac{a^i}{i!} \right)^{-1}$$

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Equilibrium distribution (2)

- Thus, the equilibrium distribution is a **truncated Poisson distribution**:

$$P\{X = i\} = \pi_i = \frac{\frac{a^i}{i!}}{\sum_{j=0}^n \frac{a^j}{j!}}, \quad i = 0, 1, \dots, n$$

- Remark (insensitivity):
 - The result is **insensitive** to the service time distribution, that is: it is valid for **any** service time distribution with mean $1/\mu$
 - So, instead of the M/M/n/n model, we can consider, as well, the more general M/G/n/n model

Time blocking

- **Time blocking** B_t = probability that all n servers are occupied at an arbitrary time = the fraction of time that all n servers are occupied
- For a stationary Markov process, this equals the probability π_n of the equilibrium distribution π . Thus,

$$B_t := P\{X = n\} = \pi_n = \frac{\frac{a^n}{n!}}{\sum_{j=0}^n \frac{a^j}{j!}}$$

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Call blocking

- **Call blocking** B_c = probability that an arriving customer finds all n servers occupied = the fraction of arriving customers that are lost
- However, due to Poisson arrivals and PASTA property, the probability that an arriving customer finds all n servers occupied equals the probability that all n servers are occupied at an arbitrary time,
- In other words, call blocking B_c equals time blocking B_t :

$$B_c = B_t = \frac{\frac{a^n}{n!}}{\sum_{j=0}^n \frac{a^j}{j!}}$$

- This is **Erlang's blocking formula**

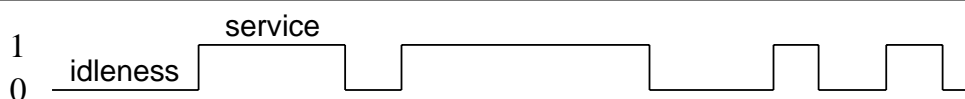
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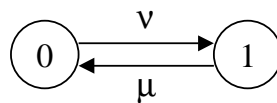
Binomial model ($M/M/k/k/k$)

- **Definition: Binomial model** is the following (simple) teletraffic model:
 - Finite number of independent customers ($k < \infty$)
 - **on-off type** customers (alternating between idleness and activity)
 - Idle times are IID and exponentially distributed with mean $1/\nu$
 - As many servers as customers ($n = k$)
 - Service times are IID and exponentially distributed with mean $1/\mu$
 - No waiting places ($m = 0$)
- Binomial model:
 - Using Kendall's notation, this is an $M/M/k/k/k$ queue
 - Although a finite system, this is clearly **lossless**
- On-off type customer (cf. lecture 4 slide 7):



On-off type customer (1)

- Let $X_j(t)$ denote the state of customer j ($j = 1, 2, \dots, k$) at time t
 - State 0 = idle, state 1 = active = in service
 - Consider what happens during a short time interval $(t, t+h]$:
 - if $X_j(t) = 0$, then, with prob. $\nu h + o(h)$, the customer becomes active (state transition $0 \rightarrow 1$)
 - if $X_j(t) = 1$, then, with prob. $\mu h + o(h)$, the customer becomes idle (state transition $1 \rightarrow 0$)
- Process $X_j(t)$ is clearly a Markov process with state transition diagram



- Note that process $X_j(t)$ is an irreducible birth-death process with a finite state space $S = \{0, 1\}$

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On-off type customer (2)

- Local balance equations (LBE):

$$\pi_0^{(j)} \nu = \pi_1^{(j)} \mu \implies \pi_1^{(j)} = \frac{\nu}{\mu} \pi_0^{(j)}$$

- Normalizing condition (N):

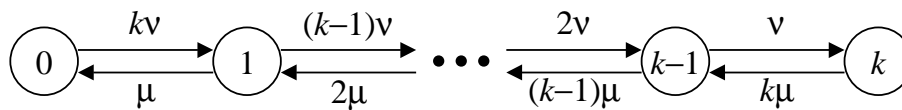
$$\pi_0^{(j)} + \pi_1^{(j)} = \pi_0^{(j)} \left(1 + \frac{\nu}{\mu}\right) = 1 \implies \pi_0^{(j)} = \frac{\mu}{\nu + \mu}, \quad \pi_1^{(j)} = \frac{\nu}{\nu + \mu}$$

- So, the equilibrium distribution of a single customer is the **Bernoulli distribution** with success probability $\nu/(\nu+\mu)$
- From this, we could deduce that the equilibrium distribution of the state of the whole system (that is: the number of active customers) is the binomial distribution $\text{Bin}(k, \nu/(\nu+\mu))$

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State transition diagram

- Let $X(t)$ denote the number of active customers
 - Assume that $X(t) = i$ at some time t , and consider what happens during a short time interval $(t, t+h]$:
 - if $i < k$, then, with prob. $(k-i)v h + o(h)$, an idle customer becomes active (state transition $i \rightarrow i+1$)
 - if $i > 0$, then, with prob. $i\mu h + o(h)$, an active customer becomes idle (state transition $i \rightarrow i-1$)
- Process $X(t)$ is clearly a Markov process with state transition diagram



- Note that process $X(t)$ is an irreducible birth-death process with a finite state space $S = \{0, 1, \dots, k\}$

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Equilibrium distribution (1)

- Local balance equations (LBE):

$$\pi_i (k-i)v = \pi_{i+1} (i+1)\mu \quad (\text{LBE})$$

$$\Rightarrow \pi_{i+1} = \frac{(k-i)v}{(i+1)\mu} \pi_i$$

$$\Rightarrow \pi_i = \frac{k!}{i!(k-i)!} \left(\frac{v}{\mu}\right)^i \pi_0 = \binom{k}{i} \left(\frac{v}{\mu}\right)^i \pi_0, \quad i = 0, 1, \dots, k$$

- Normalizing condition (N):

$$\sum_{i=0}^k \pi_i = \pi_0 \sum_{i=0}^k \binom{k}{i} \left(\frac{v}{\mu}\right)^i = 1 \quad (\text{N})$$

$$\Rightarrow \pi_0 = \left(\sum_{i=0}^k \binom{k}{i} \left(\frac{v}{\mu}\right)^i \right)^{-1} = \left(1 + \frac{v}{\mu}\right)^{-k} = \left(\frac{\mu}{v+\mu}\right)^k$$

Equilibrium distribution (2)

- Thus, the equilibrium distribution is a **binomial distribution**:

$$X \sim \text{Bin}\left(k, \frac{\nu}{\nu + \mu}\right)$$

$$P\{X = i\} = \pi_i = \binom{k}{i} \left(\frac{\nu}{\nu + \mu}\right)^i \left(\frac{\mu}{\nu + \mu}\right)^{k-i}, \quad i = 0, 1, \dots, k$$

$$E[X] = \frac{k\nu}{\nu + \mu}, \quad D^2[X] = k \cdot \frac{\nu}{\nu + \mu} \cdot \frac{\mu}{\nu + \mu} = \frac{k\nu\mu}{(\nu + \mu)^2}$$

- Remark (insensitivity):
 - The result is **insensitive both** to the service **and** the idle time distribution, that is: it is valid for **any** service time distribution with mean $1/\mu$ and **any** idle time distribution with mean $1/\nu$
 - So, instead of the $M/M/k/k/k$ model, we can consider, as well, the more general $G/G/k/k/k$ model

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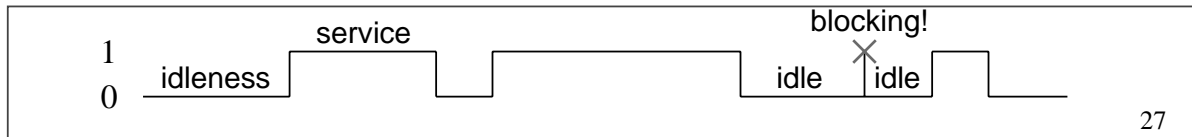
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Engset model ($M/M/n/n/k$)

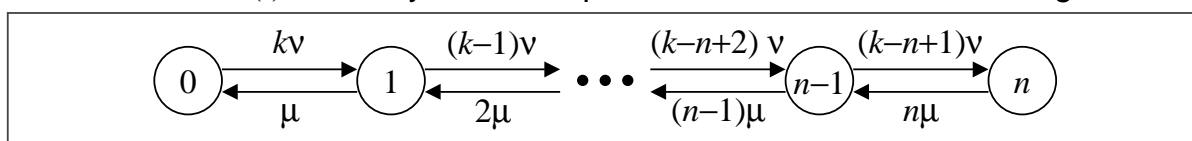
- **Definition: Engset model** is the following (simple) teletraffic model:
 - Finite number of independent customers ($k < \infty$)
 - **on-off type** customers (alternating between idleness and activity)
 - Idle times are IID and exponentially distributed with mean $1/v$
 - Less servers than customers ($n < k$)
 - Service times are IID and exponentially distributed with mean $1/\mu$
 - No waiting places ($m = 0$)
- Engset model:
 - Using Kendall's notation, this is an $M/M/n/n/k$ queue
 - This is a pure loss system, and, thus, **lossy**
- On-off type customer:

Note: If the system is full when an idle cust. tries to become an active cust., a new idle period starts.



State transition diagram

- Let $X(t)$ denote the number of active customers
 - Assume that $X(t) = i$ at some time t , and consider what happens during a short time interval $(t, t+h]$:
 - if $i < n$, then, with prob. $(k-i)v h + o(h)$, an idle customer becomes active (state transition $i \rightarrow i+1$)
 - if $i > 0$, then, with prob. $i\mu h + o(h)$, an active customer becomes idle (state transition $i \rightarrow i-1$)
- Process $X(t)$ is clearly a Markov process with state transition diagram



- Note that process $X(t)$ is an irreducible birth-death process with a finite state space $S = \{0, 1, \dots, n\}$

Equilibrium distribution (1)

- Local balance equations (LBE):

$$\pi_i(k-i)v = \pi_{i+1}(i+1)\mu \quad (\text{LBE})$$

$$\Rightarrow \pi_{i+1} = \frac{(k-i)v}{(i+1)\mu} \pi_i$$

$$\Rightarrow \pi_i = \frac{k!}{i!(k-i)!} \left(\frac{v}{\mu}\right)^i \pi_0 = \binom{k}{i} \left(\frac{v}{\mu}\right)^i \pi_0, \quad i = 0, 1, \dots, n$$

- Normalizing condition (N):

$$\sum_{i=0}^n \pi_i = \pi_0 \sum_{i=0}^n \binom{k}{i} \left(\frac{v}{\mu}\right)^i = 1 \quad (\text{N})$$

$$\Rightarrow \pi_0 = \left(\sum_{i=0}^n \binom{k}{i} \left(\frac{v}{\mu}\right)^i \right)^{-1}$$

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Equilibrium distribution (2)

- Thus, the equilibrium distribution is a **truncated binomial distribution**:

$$P\{X = i\} = \pi_i = \frac{\binom{k}{i} \left(\frac{v}{\mu}\right)^i}{\sum_{j=0}^n \binom{k}{j} \left(\frac{v}{\mu}\right)^j} = \frac{\binom{k}{i} \left(\frac{v}{v+\mu}\right)^i \left(\frac{\mu}{v+\mu}\right)^{k-i}}{\sum_{j=0}^n \binom{k}{j} \left(\frac{v}{v+\mu}\right)^j \left(\frac{\mu}{v+\mu}\right)^{k-j}}, \quad i = 0, 1, \dots, n$$

- Remark (insensitivity):

- The result is **insensitive both** to the service **and** the idle time distribution, that is: it is valid for **any** service time distribution with mean $1/\mu$ and **any** idle time distribution with mean $1/v$
- So, instead of the $M/M/n/n/k$ model, we can consider, as well, the more general $G/G/n/n/k$ model

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Time blocking

- **Time blocking** B_t = probability that all n servers are occupied at an arbitrary time = the fraction of time that all n servers are occupied
- For a stationary Markov process, this equals the probability π_n of the equilibrium distribution π . Thus,

$$B_t := P\{X = n\} = \pi_n = \frac{\binom{k}{n} \left(\frac{\nu}{\mu}\right)^n}{\sum_{j=0}^n \binom{k}{j} \left(\frac{\nu}{\mu}\right)^j}$$

Call blocking (1)

- **Call blocking** B_c = probability that an arriving customer finds all n servers occupied = the fraction of arriving customers that are lost
- In the Engset model, however, the “arrivals” do **not** follow a Poisson process. Thus, we cannot utilize the PASTA property any more.
- In fact, the distribution of the state that an “arriving” customer sees differs from the equilibrium distribution. Thus, call blocking B_c does **not** equal time blocking B_t in the Engset model.

Call blocking (2)

- Let π_i^* denote the probability that there are i active customers when an idle customer becomes active (which is called an “arrival”)
- Consider a long time interval $(0, T)$:
 - During this interval, the average time spent in state i is $\pi_i T$
 - During this time, the average number of “arriving” customers (who all see the system to be in state i) is $(k-i)v \cdot \pi_i T$
 - During the whole interval, the average number of “arriving” customers is $\sum_j (k-j)v \cdot \pi_j T$
- Thus,

$$\pi_i^* = \frac{(k-i)v \cdot \pi_i T}{\sum_{j=0}^n (k-j)v \cdot \pi_j T} = \frac{(k-i)v \cdot \pi_i}{\sum_{j=0}^n (k-j)v \cdot \pi_j}, \quad i = 0, 1, \dots, n$$

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Call blocking (3)

- It can be shown (exercise!) that

$$\pi_i^* = \frac{\binom{k-1}{i} \left(\frac{v}{\mu}\right)^i}{\sum_{j=0}^n \binom{k-1}{j} \left(\frac{v}{\mu}\right)^j}, \quad i = 0, 1, \dots, n$$

- If we write explicitly the dependence of these probabilities on the total number of customers, we get the following result:

$$\pi_i^*(k) = \pi_i(k-1), \quad i = 0, 1, \dots, n$$

- In other words, an “arriving” customer sees such a system where there is one customer less (itself!) in equilibrium

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Call blocking (4)

- By choosing $i = n$, we get the following formula for the call blocking probability:

$$B_c(k) = \pi_n^*(k) = \pi_n(k-1) = B_t(k-1)$$

- Thus, for the Engset model, the call blocking in a system with k customers equals the time blocking in a system with $k-1$ customers:

$$B_c(k) = B_t(k-1) = \frac{\binom{k-1}{n} \left(\frac{\nu}{\mu}\right)^n}{\sum_{j=0}^n \binom{k-1}{j} \left(\frac{\nu}{\mu}\right)^j}$$

- This is **Engset's blocking formula**

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THE END



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